Population pharmacokinetic analysis of elvitegravir and cobicistat in HIV-1-infected individuals

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Objectives: Co-formulated elvitegravir, cobicistat, tenofovir disoproxil fumarate and emtricitabine is among the preferred regimens for first-line ART. A population approach was used to characterize the pharmacokinetics of elvitegravir and cobicistat and identify individual factors and co-medications influencing their disposition, taking into consideration the interaction between the two compounds.

Methods: The study population included 144 HIV-infected individuals who provided 186 and 167 elvitegravir and cobicistat plasma concentrations, respectively. First, distinct NONMEM analyses were conducted for elvitegravir and cobicistat, including individual demographic, clinical and genetic factors as potential covariates. Elvitegravir and cobicistat interaction was then assessed through different inhibitory models. Simulations based on the final model served to compare expected drug concentrations under standard and alternative dosage regimens.

Results: Clearance with between-subject variability was 7.6 L/h (coefficient of variation (CV) 16.6%) and volume of distribution 61 L for elvitegravir and 16.0 L/h (CV 41.9%) and 88.3 L, respectively, for cobicistat. Concomitant administration of non-ritonavir-boosted atazanavir decreased elvitegravir clearance by 35%, likely due to UDP-glucuronosyl transferase (UGT) 1A1 inhibition. Concomitant administration of non-ritonavir-boosted atazanavir and ritonavir-boosted darunavir decreased cobicistat clearance by 47% and 27%, respectively. The final interaction model included cobicistat exposure (AUC0–2 4) on elvitegravir clearance. Simulations confirmed that a reduced elvitegravir dose of 85 mg co-administered with cobicistat and atazanavir produces a concentration–time course comparable to the standard regimen without atazanavir.

Conclusions: Elvitegravir and cobicistat pharmacokinetic variability appears to be mainly explained by drug–drug interactions that may be encountered in routine clinical practice. In these cases, therapeutic drug monitoring and surveillance for potential toxicities would be justified.

Introduction

The once-daily fixed-dose combination containing elvitegravir boosted with cobicistat, tenofovir disoproxil fumarate and emtricitabine is among the preferred regimens for first-line ART and a switch option for experienced individuals according to HIV guidelines.1–8 The integrase strand transfer inhibitor elvitegravir is also available as a stand-alone agent to be co-administered with other antiretroviral agents, combined with ritonavir or cobicistat.9–10 Elvitegravir undergoes biotransformation via cytochrome P450 3A (CYP3A) and to a lesser extent via UDP-glucuronosyl transferase (UGT) 1A1 and 1A3.11–12 Cobicistat is a novel pharmacokinetic boosting agent; it is a structural analogue of ritonavir without antiretroviral activity. It has higher specificity for inhibiting CYP3A4/3A5 and is also used to boost the PIs atazanavir and darunavir.13,14 CYP3A4 and CYP2D6 are the enzymes involved in cobicistat metabolism. Cobicistat also inhibits, though weakly, CYP2D6 and efflux transporters such as P-glycoprotein, breast cancer
resistance protein (BCRP) and organic anion-transporting polypeptides (OATPs) 1B1 and 1B3.\(^{15-17}\)

Population pharmacokinetic analyses conducted with pooled data from clinical studies in healthy and HIV-infected individuals reported moderate between-subject variability [coefficient of variation (CV)] in clearance, reaching 29%–32% for cobicistat- or ritonavir-boosted elvitegravir and 53% for cobicistat.\(^{18,19}\)

Although combination of co-formulated elvitegravir and cobicistat with other antiretroviral agents, such as ritonavir-boosted or non-boosted PIs, is not recommended, they are sometimes used in routine clinical practice to treat complicated HIV infections. Their mutual influence on each other’s concentrations is unknown and there are no dosing adjustment recommendations. Therefore, the characterization of their pharmacokinetic variability and assessment of potential drug–drug interactions outside a clinical trial setting is of relevance for clinical practice. The objective of this study was to develop a population pharmacokinetic model for elvitegravir and cobicistat in a cohort of HIV-1-infected individuals, to quantify the interaction between the two compounds and to identify potential factors influencing elvitegravir and cobicistat disposition.

Methods

Study population

Elvitegravir plasma concentrations were obtained from 144 HIV-infected individuals within the frame of therapeutic drug monitoring (TDM) performed at the University Hospital Centre of Lausanne between February 2011 and February 2015. Cobicistat plasma concentrations were measured in 133 of these individuals. A median of one sample per individual (range one to five) was collected between 0.5 and 27.75 h after the last dose intake under steady-state conditions. All participants received a fixed-dose combination containing 150 mg of elvitegravir, 150 mg of cobicistat, 300 mg of tenofovir disoproxil fumarate and 200 mg of emtricitabine in a single tablet, once daily. Exclusion criteria were undetectable elvitegravir and cobicistat plasma concentrations, suggestive of non-adherence to treatment, and non-reliable time information about blood sampling or last dose intake. This study was conducted within the frame of the Swiss HIV Cohort Study (SHCS) (http://www.shcs.ch).

Analytical method

Plasma level measurements were performed in the Laboratory of Clinical Pharmacology at the University Hospital Centre of Lausanne. Plasma samples obtained from HIV-infected individuals were isolated by centrifugation and stored at \(-20^\circ\)C until batch analysis. On the day of the analysis, samples were inactivated for virus at \(60^\circ\)C for 60 min. Plasma elvitegravir and cobicistat levels were determined by LC–MS/MS after protein precipitation with acetonitrile according to our previously reported analytical method, which was adapted to include cobicistat.\(^{20}\) The method showed acceptable precision (inter-day CV 3–6.3%) and accuracy (CV 3.8–7.2%). The calibration curves of elvitegravir and cobicistat are linear, with a lower limit of quantification (LLOQ) of 50 and 5 ng/mL, respectively. The laboratory participates in an international external quality assurance programme for antiretroviral drug analysis (KKGT, Stichting Kwaliteitsbewaking Klinische Geneesmiddelenanalyse en Toxicologie, Association for Quality Assessment in TDM and Clinical Toxicology, The Hague, The Netherlands, http://www.kkgt.nl/).

Genotyping

We selected the genetic variant UGT1A1*28 (rs8175347) with a proven loss-of-function effect and minor allelic frequency >0.05.\(^{21-23}\) We successfully genotyped 56 of the 60 SHCS participants that gave genetic consent. Genotyping was performed by direct sequencing as previously described.\(^{24}\) We verified that the variant studied was in Hardy–Weinberg equilibrium.

Population pharmacokinetic analysis

Non-linear mixed-effects modelling was performed with NONMEM\(^{25}\) (version 7.1.0. ICON Development Solutions, Ellicott City, MD, USA). First, distinct pharmacokinetic models for elvitegravir and cobicistat were built based on natural log-transformed plasma concentrations. Then, a simultaneous model describing the influence of cobicistat on elvitegravir oral clearance was developed, while considering covariates potentially influencing the pharmacokinetics of each drug.

Structural models

A stepwise procedure was used to identify models that best fitted elvitegravir and cobicistat data, testing one- and two-compartment models with first and/or zero-order absorption with and without absorption lag time. Since elvitegravir and cobicistat are administered orally, apparent clearance (CL/F) and volume of distribution (V/F) were estimated (where F is the absolute oral bioavailability).

Statistical models

Exponential errors following a log-normal distribution were assumed for the description of between-subject variability of the pharmacokinetic parameters, described by the equation \(\theta = \theta_e \times e^{\omega}\), where \(\theta_e\) is the individual pharmacokinetic parameter of the jth individual, \(\theta\) is the geometric average population value and \(\omega\) is a between-subject random effect, which follows an independent, normal distribution with mean zero and variance \(\omega^2\). A combined proportional and additive error model was used to depict residual variability.

In the cobicistat model, different approaches to the management of data below the limit of quantification (BLQ) were tested, but likelihood-based methods failed to achieve successful minimization.\(^{25-28}\) We performed a sensitivity analysis between two standard methods corresponding to the exclusion of the BLQ measurements (M1) and imputation to half the LLOQ (M2).\(^{25}\) We examined whether the BLQ data treatment affected the estimates. Finally, method M5 was used to treat BLQ data and a fixed additive variance component, corresponding to half the LLOQ, was incorporated into the error model.

Covariate models

The correlation between elvitegravir and cobicistat individual pharmacokinetic estimates and demographic, clinical and genetic covariates was first explored graphically to identify potential and physiologically plausible relationships. Then, covariate analysis was performed using a stepwise insertion/deletion approach testing linear or non-linear functions as appropriate (categorical covariates coded as 0 and 1, continuous covariates centred on their median value). Missing values were imputed to the population median value for continuous covariates. To avoid a possible bias in the covariate analysis due to missing covariates, subset analyses including only individuals with available data were performed and results were compared with the original dataset. Sex, body weight, height, body surface area, co-medications and UGT1A1*28 (rs8175347) genotype were tested as covariates on elvitegravir pharmacokinetics. Genotyped individuals were classified into three groups: homozygous for the reference allele (UGT1A1*1/*1); heterozygous (UGT1A1*1/*28); and homozygous for the rare allele (UGT1A1*28/*28). Individuals with missing genotype results were treated as a fourth group. Rich and reduced models were tested to estimate the impact of the genetic variant on individual elvitegravir pharmacokinetic parameters. Rich models allowed UGT1A1*1/*1, UGT1A1*1/*28, UGT1A1*28/*28 and the missing genotype group to have
different values of the parameter. In reduced models, genotypes with similar parameter estimates were grouped in a unique variable. Sex, body weight, height, body surface area and co-medications were tested as covariates on cobicistat pharmacokinetics.

**Interaction of cobicistat exposure on elvitegravir clearance**

Interaction models were built up from the elvitegravir intermediate covariate model developed in the covariate analysis described above. First, non-competitive models integrating cobicistat (COBI) area under the concentration–time curve from 0 to 24 h (AUCCOBI) were assessed using linear, power and exponential functions, as follows:

$$\frac{CL}{EVG} = \frac{CL_{0}}{F_{EVG}} \times \left(1 - \theta_{COBI} \times \frac{AUCCOBI \text{ median}}{AUCCOBI \text{ median}}\right)$$

$$\frac{CL}{EVG} = \frac{CL_{0}}{F_{EVG}} \times \left(\frac{AUCCOBI \text{ median}}{AUCCOBI \text{ median}}\right)^{-\theta_{COBI}}$$

$$\frac{CL}{EVG} = \frac{CL_{0}}{F_{EVG}} \times \exp\left(-\theta_{COBI} \times \frac{AUCCOBI \text{ median}}{AUCCOBI \text{ median}}\right)$$

where $CL_{0}/F_{EVG}$ is the mean population $CL/F_{EVG}$, $\theta_{COBI}$ is the factor associated with the effect of cobicistat on $CL/F_{EVG}$, $AUCCOBI$ was estimated at 9.8 mg.h/L in the study population using dose/CL/EVG, where $CL/F_{EVG}$ is the apparent clearance of cobicistat in the final model.

Competitive models were then tested using linear, exponential and maximum-effect functions, as follows:

$$\frac{CL}{EVG} = \frac{CL_{0}}{F_{EVG}} \times I(t)$$

$$I(t) = 1 - (\theta_{COBI} \times C_{COBI})$$

$$I(t) = \exp(-\theta_{COBI} \times C_{COBI})$$

$$I(t) = 1 - C_{COBI} \times \theta_{\text{max}} \frac{\theta_{\text{max}}}{C_{\text{COBI}}}$$

where $I(t)$ describes an inhibitory time-dependent model driven by cobicistat concentrations ($C_{COBI}$), $\theta_{\text{max}}$ is the factor associated with the effect of $C_{COBI}$ on $CL/F_{EVG}$, $\theta_{\text{max}}$ is the maximum inhibitory effect of cobicistat and $\theta_{\text{max}}=0.05$ is the $C_{COBI}$ producing 50% of the $I_{\text{max}}$.

**Model evaluation and assessment**

The precision of the parameters was evaluated using non-parametric bootstrapping (1000 replicates) to generate 95% CIs for parameter estimates. A prediction-corrected visual predictive check (pVPC) was also carried out using the PsN toolkit and Xpose (version 4.5.3). Observed concentrations of elvitegravir and cobicistat were visually compared with the 5th, 50th and 95th percentiles of the simulated results. Figures were generated with R (Version 3.2.2, R Development Core Team, Foundation for Statistical Computing, Vienna, Austria, http://www.r-project.org/). An external validation was also performed using 48 samples collected from a separate group of 43 HIV-infected individuals treated with co-formulated elvitegravir between February and August 2015. Quantitative and categorical variables between the model-building and validation population groups were compared by means of the Mann–Whitney test and Fisher’s exact test, respectively. Population and individual concentrations of elvitegravir and cobicistat were predicted based on the final models using the NONMEM® MAXEVAL = 0 option. The predictive performance of the model was assessed in terms of bias (mean prediction error, MPE) and precision (root mean square prediction error, RMSE) and the associated 95% CI, as follows:

$$\text{MPE} = \sum_{i=1}^{N} \frac{C_{\text{pred}} - C_{\text{obs}}}{N}$$

$$\text{RMSE} = \sqrt{\frac{\sum_{i=1}^{N} (C_{\text{pred}} - C_{\text{obs}})^2}{N}}$$

where $C_{\text{pred}}$ represents population and individual predictions, $C_{\text{obs}}$ the observed concentrations of the validation group and $N$ the number of observations.

**Simulations of standard and alternative dosage regimens**

Elvitegravir/cobicistat (150/150 mg and 85/150 mg) with and without influencing co-medication were simulated with NONMEM® for 1000 individuals based on the final model estimates, including between- and within-subject variability. The median $AUCCOBI$ of the study population was used as a covariate on $CL/F_{EVG}$. The 2.5th, 50th and 97.5th percentiles were calculated for each dosage regimen and plotted on a graph. Predicted elvitegravir $C_{\text{min}}$, $C_{\text{max}}$ and $AUCCOBI$ were estimated and used to calculate changes between different dosage regimens.

**Results**

A total of 186 and 167 observations were obtained for elvitegravir and cobicistat analyses, respectively. Participant demographic and clinical characteristics of the model-building and validation groups are summarized in Table 1. Only the proportion of male individuals was significantly different between the two datasets ($P = 0.001$). Darunavir and atazanavir were the most frequent antiretroviral co-medications in all groups. No individuals treated with atazanavir received it in combination with ritonavir, likely due to cobicistat boosting. However, 8.3% and 9.3% of individuals in the model-building and validation groups, respectively, were treated with ritonavir-boosted darunavir (darunavir/r), thus adding a second boosting agent to the antiretroviral regimen.

**Structural and covariate models for elvitegravir**

A one-compartment model with first-order absorption and elimination fitted the data correctly. Models with two compartments or alternative absorption processes, including zero-order absorption...
Table 1. Demographic and clinic characteristics of the study populations

<table>
<thead>
<tr>
<th>Characteristic</th>
<th>Elvitegravir model-building group (n=144)</th>
<th>P</th>
<th>Cobicistat model-building group (n=133)</th>
<th>P</th>
<th>Model validation group (n=43)</th>
</tr>
</thead>
<tbody>
<tr>
<td>Age (years), median (range)</td>
<td>41.0 (16–73)</td>
<td>0.09</td>
<td>41.5 (16–73)</td>
<td>0.10</td>
<td>45 (22–66)</td>
</tr>
<tr>
<td>Body weight (kg), median (range)</td>
<td>72 (45.8–125)</td>
<td>0.27</td>
<td>72 (50–125)</td>
<td>0.31</td>
<td>76 (45–110)</td>
</tr>
<tr>
<td>Height (cm), median (range)</td>
<td>175 (155–196)</td>
<td>0.27</td>
<td>175 (155–196)</td>
<td>0.22</td>
<td>173 (150–190)</td>
</tr>
<tr>
<td>Body surface area (m²), b median (range)</td>
<td>1.9 (1.5–2.6)</td>
<td>0.66</td>
<td>1.9 (1.5–2.6)</td>
<td>0.68</td>
<td>1.9 (1.4–2.3)</td>
</tr>
<tr>
<td>Race, n (%)</td>
<td></td>
<td></td>
<td></td>
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</tr>
<tr>
<td>Caucasian</td>
<td>58 (40.3)</td>
<td></td>
<td>55 (41.4)</td>
<td></td>
<td></td>
</tr>
<tr>
<td>African-American</td>
<td>13 (9.0)</td>
<td></td>
<td>12 (9.0)</td>
<td></td>
<td></td>
</tr>
<tr>
<td>Hispanic</td>
<td>4 (2.8)</td>
<td></td>
<td>2 (1.5)</td>
<td></td>
<td></td>
</tr>
<tr>
<td>Asian</td>
<td>4 (2.8)</td>
<td></td>
<td>2 (1.5)</td>
<td></td>
<td></td>
</tr>
<tr>
<td>unknown</td>
<td>65 (45.1)</td>
<td></td>
<td>62 (46.6)</td>
<td></td>
<td>43 (100)</td>
</tr>
<tr>
<td>Co-medications, n (%)</td>
<td></td>
<td></td>
<td></td>
<td></td>
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</tr>
<tr>
<td>darunavir</td>
<td>15 (10.4)</td>
<td>0.37</td>
<td>15 (11.3)</td>
<td>0.25</td>
<td>2 (4.7)</td>
</tr>
<tr>
<td>darunavir/ritonavir</td>
<td>12 (8.3)</td>
<td>0.76</td>
<td>11 (8.3)</td>
<td>0.76</td>
<td>4 (9.3)</td>
</tr>
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<td>atazanavir</td>
<td>7 (4.9)</td>
<td>0.15</td>
<td>6 (4.5)</td>
<td>0.14</td>
<td>5 (11.6)</td>
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<tr>
<td>zidovudine</td>
<td>1 (0.8)</td>
<td></td>
<td>1 (0.8)</td>
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<tr>
<td>etravirine</td>
<td>4 (2.8)</td>
<td></td>
<td>3 (2.3)</td>
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<tr>
<td>efavirenz</td>
<td>6 (4.2)</td>
<td></td>
<td>5 (3.8)</td>
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<td></td>
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<tr>
<td>nevirapine</td>
<td>1 (0.7)</td>
<td></td>
<td>1 (0.8)</td>
<td></td>
<td></td>
</tr>
<tr>
<td>maraviroc</td>
<td>3 (2.1)</td>
<td></td>
<td>3 (2.3)</td>
<td></td>
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</tbody>
</table>

P, comparison between variables in the model-building and validation population groups.

*In the model validation group the same number (n=48) of elvitegravir and cobicistat plasma concentrations were obtained.*

*bCalculated with the Mosteller formula.*

and lag time or the combination of first and zero-order absorption models with and without lag time, did not improve the fit (ΔOFV > -0.0). Between-subject variability was assigned to CL/F_EVGR with no further improvement upon its assignment to the elvitegravir apparent volume of distribution (V/F_EVGR) (ΔOFV = -0.3) or the first-order elvitegravir absorption rate constant (k_a/EVGR) (ΔOFV = -0.6). A combined proportional and additive model adequately described the residual error.

In univariate analyses, no demographic or clinical covariates tested showed any significant influence on CL/F_EVGR (ΔOFV > -3.2). The analysis of the subset of patients with complete covariate information confirmed these findings (ΔOFV > -4.1). Concomitant administration of atazanavir significantly decreased CL/F_EVGR by 54% (95% CI -7.5% to -30.5%) and reduced CL/F_EVGR variability by 12.1% (ΔOFV = -25.7, P<0.001). Co-administration of darunavir/ritonavir was associated with a decrease of CL/F_EVGR by 29.2% but this association did not reach statistical significance (95% CI -23.1% to 81.5%, ΔOFV = -2.2). Individuals in the UGT1A1*28/*28 genetic group showed a 20% decrease in CL/F_EVGR but the association was not statistically significant (95% CI -72.4% to 113%, ΔOFV = -0.9). Structural and intermediate pharmacokinetic model results for elvitegravir are summarized in Table 2.

**Structural and covariate models for cobicistat**

A one-compartment model with first-order absorption and elimination fitted the data well, whereas a two-compartment model and alternative absorption models including zero-order absorption and lag time did not improve the fit or were not retained due to the implausibility of the estimates (ΔOFV > -4.5). Between-subject variability was assigned to CL/F_COBI, with no further improvement when assigned to the cobicistat apparent volume of distribution (V/F_COBI) (ΔOFV = -13.8). Despite the observed significant fit improvement (ΔOFV = -13), the variability in the first-order absorption rate constant for cobicistat (k_a/COBI) was not retained due to the poor precision of the estimate. Six observations of cobicistat were below the LLOQ (5 ng/mL), with corresponding elvitegravir concentrations measured within the expected range of concentrations. Consequently, we could not exclude these measurements for possible non-adherence. A combined proportional and additive model was selected to describe the residual error.

In univariate analyses, no demographic or clinical covariates tested showed any significant influence on CL/F_COBI (ΔOFV > -1.2). The analysis of the subset of patients with complete covariate information provided very similar results (ΔOFV > -0.1). Concomitant...
administration of atazanavir and darunavir/r significantly reduced CL/ 
F_{COBI} by 47.3% (95% CI – 83.0% to –11.6%, ∆OFV = –8.0, P<0.005) and 
27.4% (95% Cl – 43.3% to –11.5%, ∆OFV = –8.3, P<0.005), respectively.

In multivariate analyses, both atazanavir and daruna-
vir/r co-administration remained as significant covariates. 
Cobicistat pharmacokinetic model results are summarized in 
Table 2.

### Elvitegravir and cobicistat interaction model

Models exploring the interaction of cobicistat exposure on CL/F_{EVG} 
are summarized in Table 3. A non-competitive influence of 
AUC_{COBI} on CL/F_{EVG} was better explained by a power function 
(Figure 1), which markedly improved the description of the data 
(∆OFV = –65.5, P<0.001 compared with the elvitegravir inter-
mediate covariate model). A 2-fold increase in AUC_{COBI} induced a 
46% reduction in CL/F_{EVG} and AUC_{COBI} decreased CL/F_{EVG} 

- between-subject variability by 51.3%. Atazanavir coadministra-
tion was retained as a covariate after the backward deletion step (ΔOFV = 8.3, P<0.005).

In the competitive models, the influence of cobicistat concen-
tration on elvitegravir clearance was best described by a linear 
function and improved the fit as well (ΔOFV = –35.6, P<0.001 
compared with elvitegravir intermediate covariate model).

Although there were no notable differences in goodness-of-fit 
plots and pharmacokinetic parameter precision between both 
non-competitive and competitive models, the former was chosen 
based on the AIC and BIC criterion (AAIC=ΔBIC = –29.9). The 
final population pharmacokinetic parameters for elvitegravir and 
cobicistat are summarized in Table 4. Diagnostic plots for elvite-
gravir and cobicistat final models are shown in Figures S1 and 
S2 (available as Supplementary data at JAC Online).

### Model validation

Of 1000 replicates analysed during the bootstrap analysis, 8.9% 
failed to minimize successfully and were excluded. Since all boot-
strap median values were contained within the 95% CI, differing 
by <10% from the population estimates, the model was consid-
ered reliable. The pcVPC indicates that the final models described 
the data adequately, with <10% of points outside the 90% pre-
diction interval (Figure 2). The external validation analysis showed 
a non-significant bias of 5% (95% CI – 7% to 19%) with a preci-
sion of 54%, within the range of residual variability for elvitegravir 
individual predictions and of 6% (95% CI – 8% to 22%) with a 
precision of 66% for population predictions. For cobicistat, a bias of 
–17% (95% CI – 6.9% to –26%) was observed, indicating under-
predicted individual concentrations, particularly evident during 
the absorption phase, with a precision of 56%, within the range 
of residual variably. A sub-analysis of predicted concentrations 
in the post-absorption phase, i.e. 10 h after dose administration, 
abolished the bias (–11%, 95% CI 0.4% to –21%). At the popu-
lation level, a non-significant bias of –24% (95% CI 3% to –44%) 
was observed. These results indicate that the predictive perfor-
ance of the model is adequate for the disposition phase but that 
the model slightly under-predicts peak concentrations.

### Simulations of dosage regimens

The average predicted elvitegravir C_{min} after a standard 150 mg 
daily dosage regimen at steady state was 341 ng/mL (95% 
prediction interval 192.5–541.6 ng/mL). Predicted elvitegravir AUC_{0–24} , C_{max} and C_{min} increased by 10.5 mg-h/L (IQR 9.4– 
11.6 mg-h/L, average relative increase of 53.9%), 420 ng/mL (IQR 
380.6–461.2 mg-h/L, average relative increase of 35.5%) and 
369 ng/mL (IQR 317.3–423.2 mg-h/L, average relative increase of 
112.5%), respectively, under 150/150 mg of elvitegravir/cobicistat 
co-administered with atazanavir compared with the regi-
men without atazanavir. The predicted concentration–time course 
of 85/150 mg of elvitegravir/cobicistat co-administered with ataza-
navir overlapped the 150/150 mg of elvitegravir/cobicistat concen-
tration–time profile (Figure 3). Predicted elvitegravir AUC_{0–24} 
and C_{max} decreased by 2.5 mg-h/L (IQR 2.2–2.8 mg-h/L, average 
relative decrease of 12.82%) and 274 ng/mL (IQR 260.2– 
287.8 ng/mL, average relative decrease of 23.2%), respectively, 
while predicted elvitegravir C_{min} increased by 66.2 ng/mL (IQR 
63.6–67.3 ng/mL, average relative increase of 20.4%) under
Table 3. Non-competitive and competitive models used to explore the influence of cobicistat exposure on elvitegravir apparent clearance

<table>
<thead>
<tr>
<th>Function</th>
<th>Model: $\text{CL/}F \times I^a$</th>
<th>$\text{CL/}F_{\text{EVG}}$ (RSE)</th>
<th>$\theta_{\text{COBI}}$ (RSE)</th>
<th>$\Delta\text{OFV}^b$</th>
<th>$P$</th>
<th>$\Delta\text{AIC}^b$</th>
<th>$\Delta\text{BIC}^b$</th>
</tr>
</thead>
<tbody>
<tr>
<td>Linear</td>
<td>$I = \left(1 - \theta_{\text{COBI}} \times \frac{\text{AUC}<em>{\text{COBI}} - \text{median AUC}</em>{\text{COBI}}}{\text{median AUC}_{\text{COBI}}}</td>
<td></td>
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</tr>
</tbody>
</table><p>ight)$ | 7.7 (4.7) | 0.50 (10.3) | -45.1 | &lt;0.001 | -43.1 | -39.2 |
| Power$^c$         | $I = \left(\frac{\text{AUC}<em>{\text{COBI}}}{\text{median AUC}</em>{\text{COBI}}}\right)^{-\theta_{\text{COBI}}}$ | 7.6 (4.4) | 1.12 (14.9) | -65.5 | &lt;0.001 | -63.5 | -59.6 |
| Exponential       | $I = \left(1 - \theta_{\text{COBI}} \times \frac{\text{AUC}<em>{\text{COBI}} - \text{median AUC}</em>{\text{COBI}}}{\text{median AUC}<em>{\text{COBI}}}ight)$ | 7.7 (4.6) | 0.92 (17.3) | -63.1 | &lt;0.001 | -61.1 | -57.2 |
| Linear            | $I = 1 - \theta</em>{\text{COBI}} \times C_{\text{COBI}}$ | 13.6 (18.8) | 1.18 (26.4) | -35.6 | &lt;0.001 | -33.6 | -29.7 |
| Exponential       | $I = \exp(-\theta_{\text{COBI}} \times C_{\text{COBI}})$ | 13 (14.1) | 1.88 (34.7) | -32.8 | &lt;0.001 | -30.8 | -27.0 |
| Maximum effect    | $I = 1 - \frac{C_{\text{COBI}} \times C_{\text{max}}}{\theta_{\text{IC50}} + C_{\text{COBI}}}$ | 12.6 (13.6) | 0.381$^d$ (40.7) | -30.4 | &lt;0.001 | -28.4 | -24.6 |</p>

$\Delta\text{OFV}$, difference in the NONMEM$^b$ objective function value; $\Delta\text{AIC}$, difference in the Akaike information criterion; $\Delta\text{BIC}$, difference in the Bayesian information criterion or Schwarz criterion; $\text{CL/}F_{\text{EVG}}$, elvitegravir apparent clearance (L/h); $\text{AUC}_{\text{COBI}}$, cobicistat plasma concentration at each timepoint; $\theta_{\text{COBI}}$, factor associated with the effect of cobicistat pharmacokinetics on $\text{CL/}F_{\text{EVG}}$; $C_{\text{max}}$, maximum inhibitory effect of cobicistat fixed to 1 as no consistent estimate could be obtained; $\theta_{\text{IC50}}$, estimate of cobicistat concentration (ng/mL) producing 50% of the $C_{\text{max}}$; RSE, relative standard error expressed as percentage and defined as $\text{SE/estimate}$, with SE directly retrieved from NONMEM$^b$.

$^a$I describes an inhibitory model induced by cobicistat exposures or timepoint concentrations.

$^b$Compared with elvitegravir intermediate covariate model.

$^c$Non-competitive model included in elvitegravir final population pharmacokinetic model.

$^d$$\theta_{\text{IC50}}$.

Discussion

Our study provides a description of the population pharmacokinetic profile of elvitegravir and cobicistat in a real patient setting, allowing simulations of standard and alternative dosage regimens. These population pharmacokinetic models can be used to derive percentile reference curves or implemented in a Bayesian tool for dosage adjustment. The CL/F$_{\text{EVG}}$ pharmacokinetic estimate is close to reported values of 6.6–7.1 L/h, while V/F$_{\text{EVG}}$ is much lower than previously published values of 137.9–197.9 L, probably due to study design distinctions, since data were collected over a 28 h period, hindering the estimation of an extensive drug distribution in peripheral tissues. Although the elvitegravir absorption phase could not be exhaustively described, with scarce measurements available within a few hours after drug administration, a mean absorption time of about 6 h is comparable to values obtained with other more complex published models. The CL/F$_{\text{COBI}}$ and V/F$_{\text{COBI}}$ pharmacokinetic estimates are in good agreement with previously reported values of 15.0 L/h and 77.0 L, respectively. Similarly, the cobicistat mean absorption time of about 1.9 h is close to values obtained with more comprehensive absorption models. None of the demographic and clinical covariates tested appeared to have a significant impact on elvitegravir and cobicistat elimination, in contrast to results suggesting statistically significant, although not clinically relevant, influences of body surface area on CL/F$_{\text{COBI}}$ and of body weight on CL/F$_{\text{COBI}}$. This could be explained by the lack of extreme body surface area values in our study population.

85/150 mg of elvitegravir/cobicistat compared with the standard regimen without atazanavir.
Table 4. Final population pharmacokinetic parameters for elvitegravir and cobicistat with bootstrap results

<table>
<thead>
<tr>
<th>Parameter</th>
<th>Estimate</th>
<th>RSE (%)</th>
<th>Bootstrap (n=1000 samples)</th>
</tr>
</thead>
<tbody>
<tr>
<td></td>
<td></td>
<td></td>
<td>Median</td>
</tr>
<tr>
<td>Elvitegravir CL/F_{EVG} (L/h)</td>
<td>7.6</td>
<td>4.4</td>
<td>7.6</td>
</tr>
<tr>
<td>V/F_{EVG} (L)</td>
<td>61.0</td>
<td>16.7</td>
<td>62.8</td>
</tr>
<tr>
<td>k_{a_{EVG}} (h^{-1})</td>
<td>0.17</td>
<td>16.4</td>
<td>0.17</td>
</tr>
<tr>
<td>$\theta_{ATV_{EVG}}$</td>
<td>-0.35</td>
<td>25.7</td>
<td>-0.36</td>
</tr>
<tr>
<td>$\theta_{AUC_{COBI}}$</td>
<td>1.12</td>
<td>14.9</td>
<td>1.13</td>
</tr>
<tr>
<td>BSV CL/F_{EVG} (%)</td>
<td>16.6</td>
<td>31.2</td>
<td>18.3</td>
</tr>
<tr>
<td>additive residual error (ng/mL)</td>
<td>148</td>
<td>28.0</td>
<td>133</td>
</tr>
<tr>
<td>proportional residual error (%)</td>
<td>39.8</td>
<td>10.9</td>
<td>38.4</td>
</tr>
<tr>
<td>Cobicistat CL/F_{COBI} (L/h)</td>
<td>16.0</td>
<td>11.3</td>
<td>15.9</td>
</tr>
<tr>
<td>V/F_{COBI} (L)</td>
<td>88.3</td>
<td>15.2</td>
<td>84.3</td>
</tr>
<tr>
<td>k_{a_{COBI}} (h^{-1})</td>
<td>0.76</td>
<td>40.6</td>
<td>0.75</td>
</tr>
<tr>
<td>$\theta_{ATV_{COBI}}$</td>
<td>-0.47</td>
<td>38.4</td>
<td>-0.47</td>
</tr>
<tr>
<td>$\theta_{DRV}$</td>
<td>-0.27</td>
<td>30.1</td>
<td>-0.27</td>
</tr>
<tr>
<td>BSV CL/F_{COBI} (%)</td>
<td>41.9</td>
<td>12.1</td>
<td>41.5</td>
</tr>
<tr>
<td>additive residual error (ng/mL)</td>
<td>2.5^a</td>
<td>12.0</td>
<td>2.5^a</td>
</tr>
<tr>
<td>proportional residual error (%)</td>
<td>2.5^a</td>
<td></td>
<td></td>
</tr>
</tbody>
</table>

Final elvitegravir model: \(\text{CL/F}_{\text{EVG}} = 7.6 \times (1 - 0.35 \times \text{ATV}) \times (\text{AUC}_{\text{COBI}}/\text{median AUC}_{\text{COBI}})^{-1.12}\). Final cobicistat model: \(\text{CL/F}_{\text{COBI}} = 16.0 \times (1 - 0.47 \times \text{ATV}) \times (1 - 0.27 \times \text{DRV})\).

\(\text{CL/F}_{\text{EVG}}\), elvitegravir apparent clearance; \(\text{V/F}_{\text{EVG}}\), elvitegravir apparent volume of distribution; \(k_{a_{EVG}}\), first-order elvitegravir absorption rate constant; \(\theta_{ATV_{EVG}}\), relative influence of atazanavir co-administration on \(\text{CL/F}_{\text{EVG}}\); \(\theta_{AUC_{COBI}}\), relative influence of cobicistat AUC (\(\text{AUC}_{\text{COBI}}\)) on \(\text{CL/F}_{\text{EVG}}\) following a power function; \(\text{CL/F}_{\text{COBI}}\), cobicistat apparent clearance; \(\text{V/F}_{\text{COBI}}\), cobicistat apparent volume of distribution; \(k_{a_{COBI}}\), first-order cobicistat absorption rate constant; \(\theta_{ATV_{COBI}}\), relative influence of atazanavir co-administration on \(\text{CL/F}_{\text{COBI}}\); \(\theta_{DRV}\), relative influence of darunavir/r co-administration on \(\text{CL/F}_{\text{COBI}}\); BSV, estimate of between-subject variability; RSE, relative standard error defined as SE/estimate, with standard error (SE) directly retrieved from NONMEM®.

^aAdditive part of cobicistat error model was fixed to $$\sqrt{0.25 \times (\text{LLOQ})^2} = 2.5 \, \text{ng/mL}$$ to account for BLQ data.

Figure 2. pcVPC of the elvitegravir (n=186) and cobicistat (n=167) final model. Open circles represent elvitegravir (left-hand panel) and cobicistat (right-hand panel) plasma concentrations. The continuous line represents the population median prediction from the final model and the broken lines represent the 90% prediction intervals. PI, prediction interval.
The results of our study highlight relevant drug interactions likely to be encountered in routine clinical practice. The significant impact of concomitant atazanavir on CL/FEVG is likely explained by the known atazanavir-mediated inhibition of UGT1A1, an enzyme that mediates the secondary metabolic pathway of elvitegravir. The interaction between ritonavir-boosted elvitegravir and atazanavir was evaluated in the early stages of elvitegravir development, showing significantly higher elvitegravir exposures when 200/100 mg of elvitegravir/ritonavir was co-administered with 300 mg of atazanavir. Hence, a reduced dose of 85 mg of elvitegravir was finally selected for the combination with 300/100 mg of atazanavir/ritonavir and recently commercialized in some countries. Our results confirm the interaction between elvitegravir and atazanavir and validate the dose adjustment proposed for the combination of elvitegravir with atazanavir boosted with ritonavir or cobicistat. As shown in previous interaction studies, our simulations predicted a slightly higher CL/FEVG and AUC 0–24 for the adapted elvitegravir dosage regimen, with 300 mg of atazanavir. Hence, a reduced dose of 85 mg of atazanavir was finally selected for the combination with 300/100 mg of atazanavir/ritonavir and recently commercialized in some countries. Our results confirm the interaction between elvitegravir and atazanavir and validate the dose adjustment proposed for the combination of elvitegravir with atazanavir boosted with ritonavir or cobicistat. As shown in previous interaction studies, our simulations predicted a slightly higher CL/FEVG and AUC 0–24 for the adapted elvitegravir dosage regimen, probably of limited clinical relevance. The question remains whether the same elvitegravir dose adjustment is suitable with co-administration of 400 mg of atazanavir, since only one individual in the study population was under this dosage regimen.

Homozygosity for the allele UGT1A1*28 led to a mild reduction in CL/FEVG, which was statistically non-significant, probably due to the small sample size. A 20% reduction in CL/FEVG might not be clinically relevant compared with the moderate between-subject variability, but should be confirmed in larger studies.

Atazanavir and darunavir are known to inhibit CYP3A4, which mediates the main metabolic pathway of cobicistat and explains the influence of both drugs on CL/FEVG. These drug–drug interactions might also have an indirect impact on elvitegravir pharmacokinetics. The effect of atazanavir on CL/FEVG is probably explained both indirectly by the increase in cobicistat exposure and by the direct UGT1A1 inhibition. None of these covariates decreased between-subject variability of CL/FEVG, probably due to difficulties in distinguishing between-from within-subject variability in our analysis, which included only one observation for most individuals.

Published data on cobicistat-boosted elvitegravir did not show a significant effect of AUCCOBI on elvitegravir clearance. However, ritonavir AUC0–24 (dose ranging from 100 to 200 mg) had a statistically significant but clinically irrelevant effect on elvitegravir bioavailability. In our study, AUCCOBI showed a marked influence on elvitegravir elimination when modelled as a non-competitive inhibitor. Notably, this effect explained up to half of the elvitegravir between-subject variability. A reduction of AUCCOBI by 25%, 50% and 75% would drive an increase in CL/FEVG and therefore a reduction in elvitegravir exposure of 38.0%, 117.3% and 372.4%, respectively. These results emphasize the need for a careful examination of co-medications modulating CYP3A4 activity that could directly and indirectly affect elvitegravir concentrations.

Elvitegravir antiviral effect was supported by a Phase 2 study in which ritonavir-boosted elvitegravir Cmin values exceeded the protein-binding-adjusted 95% inhibitory concentration (IC95) of 45 ng/mL and were associated with viral suppression over 10 days (2 log reduction from baseline in HIV-1 RNA copies). No association between elvitegravir exposure and virological response by week 48 was found in Phase 3 pharmacokinetic–pharmacodynamic studies and no therapeutic target, apart from the IC95, has been defined to date. Model-based simulated minimal concentrations suggest that all individuals will present concentrations above the IC95 and would be at low risk of virological failure with the recommended dosage regimen. On the other hand, one might conjecture whether a proportion of individuals might be exposed to concentrations higher than the required levels without any substantial benefit, but possibly with undesirable consequences in the long term. However, there is currently no clinical evidence to support a cumulative toxicity of elvitegravir.

In conclusion, the variability in elvitegravir pharmacokinetics is modest after accounting for variations in cobicistat exposure. Drug–drug interactions that might be encountered in daily practice might lead to inadequate elvitegravir levels and monitoring of plasma concentrations could be a useful tool to identify patients at risk of suboptimal concentrations.

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Members of the Swiss HIV Cohort Study
proxil fumarate for initial treatment of HIV-1 infection: a randomised, double-blind, placebo-controlled trial versus ritonavir-boosted protease inhibitor with emtricitabine and tenofovir disoproxil fumarate (TDF/FTC) over 48 weeks. The results of this study are compared with the currently accepted treatment standard, which is a regimen of TDF/FTC boosted with ritonavir. The study aimed to determine the efficacy and safety of elvitegravir/cobicistat/emtricitabine/tenofovir (ETV/CBI/FTC/TDF) compared to TDF/FTC boosted with ritonavir (RTV) over 48 weeks in treatment-naive patients with HIV-1 infection. The study was conducted at the University of Lausanne, Switzerland, in collaboration with various Swiss hospitals and private physicians.

**Funding**

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**Supplementary data**

Figures S1 and S2 are available as Supplementary data at JAC Online (http://jac.oxfordjournals.org/).

**References**


